

Polyhedral risk measures in stochastic programming

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1 Introduction

When wishing to replace the usual **expectation**-based objective by some functional measuring **risk**, at least the following three issues have to be addressed:

- What is an **appropriate** risk measure for the underlying **practical model** ?
- Does this exchange lead to **serious changes in structure and stability** ?
- Does the exchange cause **serious computational problems** ?

References: Schultz/Tiedemann 02, Schultz 03

Of course, the stochastic programming user wishes that his choice of a risk measure leads to the answer **no** on the last two questions.

We will discuss a **subclass of convex risk measures** having this enjoyable property.

Our **motivation** stems from stochastic programming applications in **electricity portfolio management**, i.e., from solving large scale mixed-integer stochastic programs.

2 Polyhedral risk measures

Let \mathcal{Z} denote a linear space of real random variables on some probability space $(\Omega, \mathcal{F}, \mathbb{P})$. We assume that \mathcal{Z} contains the constants. A functional $\rho : \mathcal{Z} \rightarrow \mathbb{R}$ is called a **risk measure** if it satisfies the following two conditions for all $z, \tilde{z} \in \mathcal{Z}$:

- (i) If $z \leq \tilde{z}$, then $\rho(z) \geq \rho(\tilde{z})$ (**monotonicity**).
- (ii) For each $r \in \mathbb{R}$ we have $\rho(z + r) = \rho(z) - r$ (**translation invariance**).

A risk measure ρ is called **convex** if it satisfies the condition

$$\rho(\lambda z + (1 - \lambda)\tilde{z}) \leq \lambda\rho(z) + (1 - \lambda)\rho(\tilde{z})$$

for all $z, \tilde{z} \in \mathcal{Z}$ and $\lambda \in [0, 1]$.

A convex risk measure is called **coherent** if it is positively homogeneous, i.e., $\rho(\lambda z) = \lambda\rho(z)$ for all $\lambda \geq 0$ and $z \in \mathcal{Z}$.

References: Artzner/Delbaen/Eber/Heath 99, Föllmer/Schied 02

Definition: A risk measure ρ on \mathcal{Z} will be called **polyhedral** if there exist $k, l \in \mathbb{N}$, $a, c \in \mathbb{R}^k$, $q, w \in \mathbb{R}^l$, a polyhedral set $X \subseteq \mathbb{R}^k$ and a polyhedral cone $Y \subseteq \mathbb{R}^l$ such that

$$\rho(z) = \inf \{ \langle c, x \rangle + \mathbb{E}[\langle q, y \rangle] : \langle a, x \rangle + \langle w, y \rangle = z, x \in X, y \in Y \}$$

for each $z \in \mathcal{Z}$. Here, \mathbb{E} denotes the expectation on $(\Omega, \mathcal{F}, \mathbb{P})$ and $\langle \cdot, \cdot \rangle$ the scalar product on \mathbb{R}^k .

The notion *polyhedral* risk measure is motivated by the polyhedrality of $\rho(z)$ as a function of the scenarios of z if z is discrete.

Assume that ρ is a polyhedral risk measure on the space $\mathcal{Z} = L_1(\Omega, \mathcal{F}, \mathbb{P})$, that $\langle w, Y \rangle = \mathbb{R}$ and $\{u \in \mathbb{R} : uw - q \in Y^*\} \neq \emptyset$, where Y^* is the polar cone of Y . Then there exist two real numbers u_ℓ , $\ell = 1, 2$, such that

$$\rho(z) = \inf_{x \in X} \{ \langle c, x \rangle + \mathbb{E} [\max_{\ell=1,2} u_\ell (z - \langle a, x \rangle)] \}.$$

In particular, ρ is a *convex* risk measure. It is *coherent* if X is a cone.

Proposition: Let $k, l \in \mathbb{N}$, $a, c \in \mathbb{R}^k$, $q, w \in \mathbb{R}^l$, a polyhedral set $X \subseteq \mathbb{R}^k$ and a polyhedral cone $Y \subseteq \mathbb{R}^l$ be given such that

$$\rho(z) = \inf \{ \langle c, x \rangle + \mathbb{E}[\langle q, y \rangle] : \langle a, x \rangle + \langle w, y \rangle = z, x \in X, y \in Y \}$$

for each $z \in L_1(\Omega, \mathcal{F}, \mathbb{P})$.

Let $\langle w, Y \rangle = \mathbb{R}$ and $\emptyset \neq \{u \in \mathbb{R} : uw - q \in Y^*\} \subset \mathbb{R}_-$ and a, c and X have the form $a = (\hat{a}, -1)$, $c = (\hat{c}, 1)$ and $X = \hat{X} \times \mathbb{R}$, where $\hat{a}, \hat{c} \in \mathbb{R}^{k-1}$ and $\hat{X} \subseteq \mathbb{R}^{k-1}$.

Then ρ is a **convex risk measure** on $L_1(\Omega, \mathcal{F}, \mathbb{P})$ if it is finite.

Furthermore, if ρ is finite it admits the following **dual representation**

$$\rho(z) = \sup \left\{ -\mathbb{E}[\lambda z] + \inf_{\hat{x} \in \hat{X}} \langle \hat{c} + \hat{a}, \hat{x} \rangle : \lambda \in L_{p'}(\Omega, \mathcal{F}, \mathbb{P}), \mathbb{E}[\lambda] = 1, -(q + \lambda w) \in Y^* \right\}$$

for each $z \in L_p(\Omega, \mathcal{F}, \mathbb{P})$ with $1 < p < +\infty$ and $\frac{1}{p} + \frac{1}{p'} = 1$.

Proof: by relying on stochastic programming methodology; the Lagrangian dual function has the form $D(\lambda) = \inf_{x \in X} \{ \langle c + \mathbb{E}[\lambda]a, x \rangle + \inf_{y \in L_p, y \in Y} \mathbb{E}[\langle q + \lambda w, y \rangle] - \mathbb{E}[\lambda z] \}$ (Reference: Wets 70).

Example 1: Let $k = l = 1$, $a = -1$, $c = 1$ and $X = \mathbb{R}$. Then the conditions in the Proposition imply $Y = \mathbb{R}$, $w \neq 0$ and $\frac{q}{w} < 0$. We obtain that ρ has the form

$$\rho(z) = \mathbb{E}\left[\frac{q}{w}z\right] + \inf\left\{\left(1 + \frac{q}{w}\right)x : x \in \mathbb{R}\right\}.$$

Hence, ρ is finite iff $\frac{q}{w} = -1$ iff $\rho(z) = \mathbb{E}[-z]$.

Example 2: [Conditional](#), [Tail](#) or [Average Value at Risk](#)

We consider the *Average Value at Risk* $AVaR_\alpha$ defined by

$$\begin{aligned} AVaR_\alpha(z) &:= \frac{1}{\alpha} \int_0^\alpha VaR_\gamma(z) d\gamma \\ &= \inf_{r \in \mathbb{R}} \left\{ r + \frac{1}{\alpha} \mathbb{E}[\max\{0, -r - z\}] \right\}, \end{aligned}$$

where $VaR_\alpha(z) := \inf\{r \in \mathbb{R} : \mathbb{P}(z + r < 0) \leq \alpha\}$ is the Value at Risk at level $\alpha \in (0, 1)$.

$AVaR_\alpha$ is polyhedral by setting $k = 1$, $l = 2$, $a = -1$, $c = 1$, $q = (\frac{1}{\alpha}, 0)$, $w = (-1, 1)$, $X = \mathbb{R}$ and $Y = \mathbb{R}_+^2$.

The condition $-(q + \lambda w) \in Y^*$ in the dual representation is equivalent to $\lambda \in [0, \frac{1}{\alpha}]$.

3 Multiperiod polyhedral risk measures

While the notion of a polyhedral risk measure is appropriate for two-stage stochastic programming models, we now consider a multiperiod extension in case that instead of the real random variable z a real stochastic process $z = \{z_t, \mathcal{F}_t\}_{t=2}^T$ with a filtration $\mathcal{F}_2 \subseteq \dots \subseteq \mathcal{F}_t \subseteq \dots \subseteq \mathcal{F}_T = \mathcal{F}$ is given. It is assumed that z_τ , $\tau = 2, \dots, t$, is measurable with respect to \mathcal{F}_t , $t = 2, \dots, T$.

As natural candidates for such an extension we consider functionals ρ on the linear space $\times_{t=2}^T L_1(\Omega, \mathcal{F}_t, \mathbb{P})$ that are defined as optimal values of specific **multi-stage stochastic programs**. Namely, we assume that there are $k_t \in \mathbb{N}$, $c_t \in \mathbb{R}^{k_t}$, $t = 1, \dots, T$, $a_t \in \mathbb{R}^{k_{t-1}}$, $w_t \in \mathbb{R}^{k_t}$, $t = 2, \dots, T$, a polyhedral set $Y_1 \subseteq \mathbb{R}^{k_1}$ and polyhedral cones $Y_t \subseteq \mathbb{R}^{k_t}$, $t = 2, \dots, T$, such that

$$\rho(z) = \inf \left\{ \mathbb{E} \left[\sum_{t=1}^T \langle c_t, y_t \rangle \right] : y_1 \in Y_1, y_t \text{ is } \mathcal{F}_t\text{-measurable, } y_t \in Y_t, \right. \\ \left. \langle a_t, y_{t-1} \rangle + \langle w_t, y_t \rangle = z_t, t = 2, \dots, T \right\}$$

Dynamic programming formulation of ρ :

$$\begin{aligned}\rho(z) &= \inf\{\langle c_1, y_1 \rangle + \mathbb{E}[v_2(y_1, z)] : y_1 \in Y_1\} \\ v_t(y_{t-1}, z_t, \dots, z_T) &:= \inf\{\langle c_t, y_t \rangle + \mathbb{E}[v_{t+1}(y_t, z_{t+1}, \dots, z_T) | \mathcal{F}_t] : \\ &\quad y_t \in Y_t, \langle a_t, y_{t-1} \rangle + \langle w_t, y_t \rangle = z_t\}, \\ &\quad t = T, \dots, 2, \\ v_{T+1}(y_T) &:= 0.\end{aligned}$$

References: Rockafellar/Wets 76, Evstigneev 76

Dual formulation for ρ :

$$\begin{aligned}\rho(z) &= \sup\left\{ \inf_{y_1 \in Y_1} \langle c_1 + \mathbb{E}[\lambda_2] a_2, y_1 \rangle - \mathbb{E} \left[\sum_{t=2}^T \lambda_t z_t \right] : \right. \\ &\quad \lambda_t \in L_{p'}(\Omega, \mathcal{F}_t, \mathbb{P}), t = 2, \dots, T, -(c_T + \lambda_T w_T) \in Y_T^*, \\ &\quad \left. -(c_t + w_t \lambda_t + a_{t+1} \mathbb{E}[\lambda_{t+1} | \mathcal{F}_t]) \in Y_t^*, t = T-1, \dots, 2 \right\}\end{aligned}$$

holds whenever $z_t \in L_p(\Omega, \mathcal{F}_t, \mathbb{P})$, $t = 2, \dots, T$, $p > 1$,
 $\frac{1}{p} + \frac{1}{p'} = 1$, and the right-hand side is finite.

References: Eisner/Olsen 75, Rockafellar/Wets 76, 78

Definition: (Artzner/Delbaen/Eber/Heath/Ku 02)

A functional ρ on (a subset of) $\times_{t=2}^T L_p(\Omega, \mathcal{F}_t, \mathbb{P})$ ($p > 1$) is called a **multi-period coherent** risk measure if it satisfies the Fatou property and if there exist positive reals $\eta_t > 0$, $t = 2, \dots, T$, $\sum_{t=2}^T \eta_t = 1$, and a closed convex set Λ of nonnegative functions contained in $\times_{t=2}^T L_{p'}(\Omega, \mathcal{F}_t, \mathbb{P})$, with $\frac{1}{p} + \frac{1}{p'} = 1$, such that

$$\rho(z) = \sup \left\{ - \sum_{t=2}^T \eta_t \mathbb{E}[\lambda_t z_t] : \lambda \in \Lambda, \sum_{t=2}^T \eta_t \mathbb{E}[\lambda_t] = 1 \right\}.$$

$$\begin{aligned} \Lambda(\eta) := & \left\{ \lambda \in \times_{t=2}^T L_{p'}(\Omega, \mathcal{F}_t, \mathbb{P}) : -\left(c_T + \lambda_T \frac{w_T}{\eta_T}\right) \in Y_T^*, \right. \\ & \left. -\left(c_t + \frac{w_t}{\eta_t} \lambda_t + \frac{a_{t+1}}{\eta_{t+1}} \mathbb{E}[\lambda_{t+1} | \mathcal{F}_t]\right) \in Y_t^*, t = 2, \dots, T-1 \right\} \end{aligned}$$

Proposition:

A multi-period polyhedral risk measure is multi-period coherent if

- (i) $\eta_t = \frac{1}{T-1}$, $t = 2, \dots, T$,
- (ii) $\lambda \in \Lambda(\eta)$ implies $\lambda_t \geq 0$, \mathbb{P} -a.s., $t = 2, \dots, T$, and
- (iii) $\lambda \in \Lambda(\eta)$ and $\inf_{y_1 \in Y_1} \langle c_1 + \mathbb{E}[\lambda_2] a_2, y_1 \rangle = 0$ imply $\mathbb{E}[\lambda_t] = 1$, $t = 2, \dots, T$.

4 Multistage SP models with minimal risk

We consider a stochastic process $\{\xi^t, \mathcal{F}_t\}_{t=2}^T$ with distribution P modelling the uncertain future and the multi-stage stochastic program

$$\text{Minimize} \quad \sum_{i=1}^I C_{i1}x_i^1 + \mathbb{E} \left[\sum_{t=2}^T \sum_{i=1}^I C_{it}(\xi^t)x_i^t \right] \quad (\text{expectation})$$

or, alternatively,

$$\text{Minimize} \quad \sum_{i=1}^I C_{i1}x_i^1 + \rho \left(\left\{ \sum_{i=1}^I C_{it}(\xi^t)x_i^t \right\}_{t=2}^T \right) \quad (\text{risk measure})$$

such that

$$\begin{aligned} x^t \text{ is } \mathcal{F}_t \text{ - measurable, } x_i^t &\in X_{it}, t = 1, \dots, T, \\ A_{it,t}x_i^t + A_{it,t-1}(\xi^t)x_i^{t-1} &\geq g_{it}(\xi^t), t = 2, \dots, T, i = 1, \dots, I, \\ \sum_{i=1}^I B_{it}(\xi^t)x_i^t &\geq d_t(\xi^t), t = 1, \dots, T. \end{aligned}$$

Here, $x^t = (x_1^t, \dots, x_I^t)$, $t = 1, \dots, T$, $\mathcal{F}_1 = \{\emptyset, \Omega\}$, $\mathcal{F}_T = \mathcal{F}$, $A_{it,\tau}$, $\tau = t-1, t$, B_{it} , g_{it} and d_t are matrices and vectors possibly depending on ξ^t , $t = 1, \dots, T$, and X_{it} subsets of Euclidean spaces.

Incorporating the multiperiod risk measure into the original program leads to a multistage model exhibiting the same structure and having (x, y) as decision.

$$\min_{(x,y)} \sum_{i=1}^I C_{i1} x_i^1 + \mathbb{E} \left[\sum_{t=1}^T \langle c_t, y_t \rangle \right]$$

such that

$$y_1 \in Y_1, y_t \text{ is nonanticipative}, y_t \in Y_t,$$

$$\langle a_t, y_{t-1} \rangle + \langle w_t, y_t \rangle = \sum_{i=1}^I C_{it}(\xi^t) x_i^t, t = 2, \dots, T,$$

$$x^t \text{ is nonanticipative}, x_i^t \in X_{it}, t = 2, \dots, T,$$

$$A_{it,t} x_i^t + A_{it,t-1}(\xi^t) x_i^{t-1} \geq g_{it}(\xi^t), t = 2, \dots, T, i = 1, \dots, I,$$

$$\sum_{i=1}^I B_{it}(\xi^t) x_i^t \geq d_t(\xi^t), t = 1, \dots, T.$$

The stability behaviour and the metrics μ_c as well as the decomposition structure (e.g. for scenario, node and geographic decomposition) do (almost) not change !